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Comparisons of Cutoff and Regression-Based Definitions of Reading Disabilities

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This study addressed the issue of specificity in reading disability by comparing two approaches to defining and selecting children with reading disabilities. One approach defined reading disability according to cutoff scores representing appropriate levels of intelligence and reading deficiency, whereas the other approach adjusted these scores for their intercorrelation through regression procedures. Results revealed clear differences in which children were identified as reading disabled according to the two definitions. However, differences in neuropsychological performance between children whose reading scores were discrepant or not discrepant with IQ were small and nonspecific for both definitions. The results of this study show that children identified as reading disabled vary according to the definition employed; at this point, there is little evidence suggesting any specificity of reading disability according to definition.

Careful analyses of the criteria used to define and select children with reading disabilities are urgently needed. There is a persistent tendency on the part of researchers and professionals working with this population to accept traditional definitions of reading disability based on consensus or professional opinion without examination of the assumptions or empirical characteristics of various definitions (Fletcher & Morris, 1986). It is most common to define a child as reading disabled in the presence of a discrepancy between potential ability and the actual achievement of reading skills. Terms such as "dyslexia" and "specific reading disability" have been used almost interchangeably to refer to children with average levels of potential ability, poor reading skills, and an absence of problems due to cultural factors, instructional methods, and sensory or acquired neurological deficits. However, neither the labels nor the criteria for their operationalization have been subjected to extensive analyses of validity and reliability (Fletcher & Morris, 1986; Rourke, 1983; Rutter, 1974). As Rourke (1983) suggested, these definitions have rarely emerged from empirical investigation.

The analysis of these criteria, a common enterprise in classification research,

is necessary for any scientific endeavor addressed to subpopulations presumably sharing common phenotypic characteristics (Fletcher, Francis, & Morris, 1988; Morris & Fletcher, 1988; Rourke, 1983). Classification research should address the validity of definitions that are often implicit in describing a child as reading disabled. These definitions should be based on criteria that are explicit, public, and clearly operationalized. When criteria vary, subject characteristics change across studies, leading to intrasubject variability that obscures attempts at measurement of skills or studies of remediation (Doehring, 1978). Similarly, classification is the basis for a common nomenclature that must be explicit for accurate communication among professionals. Finally, it is difficult to interpret experimental research without careful specification of the independent variables. Regardless of how well a construct is measured (i.e., dependent variables), studies that don't specify carefully operationalized criteria for selecting subjects cannot be replicated and null results become difficult to interpret (Fletcher et al., 1988). Research on children with reading disabilities highlights the need for classification studies addressing the validity of commonly accepted definitions used

to identify children as reading disabled (Fletcher & Morris, 1986; Morris & Fletcher, 1988; Rutter, 1974).

An example of these problems is the use of discrepancy-based definitions of reading disability. When children with reading problems are defined and selected according to discrepancies between intelligence and achievement—a common practice since the adoption of Public Law 94-142—several problems emerge. One set of problems concerns the use of IQ tests to measure potential. A child's performance on IQ measures is influenced by past learning, genetic endowment, and a host of situational factors (Bortner & Birch, 1980; Estes, 1981). In addition, the same cognitive problems that are present in children with reading problems also reduce scores on IQ tests. For example, a processing deficiency in language, which may be associated with reading problems, may also reduce verbal IQ scores (Taylor, Fletcher, & Satz, 1984). Finally, IQ tests such as the Wechsler Intelligence Scale for Children—Revised (WISC-R) do not measure all aspects of a child's adaptive functioning, such as social behavior and more specific information processing skills (Guilford, 1967; Kaufman, 1979). Most IQ scores represent combined measurements of several covarying abilities—hardly a measure of a single construct called "potential." This is illustrated by the fact that IQ scores are correlated with school achievement in randomly selected groups of children, but have not been shown to be particularly predictive of achievement in children with reading disabilities (Sattler, 1982; Taylor et al., 1984).

The second set of problems is statistical and concerns the attempt to define learning disabilities by statistical comparisons of IQ and achievement standard scores. Such comparisons are apparent in eligibility legislation in most states and school districts across the United States. The statistical comparison of discrepancies between imperfectly correlated test scores is always associated with regression to the mean. Basically, regression to the mean results from the interdependence of imperfectly correlated measures. Practically, if scores on an IQ test and an achievement test are not corrected for their intercorrelation, there will be a tendency for scores on one test to move

closer to an average score in the presence of an extreme score on the other test. For example, children with IQ scores 2 standard deviations above the mean will, *on the average*, have achievement scores less than 2 standard deviations above the mean (Yule, 1978).

It is possible to use regression procedures to define IQ-achievement discrepancies that correct for regression artifacts (Reynolds, 1984). However, even when scores are corrected, it is not clear that children with discrepancies in IQ and achievement have more specific disabilities than do poor achievers whose IQ scores are not discrepant. Rutter and Yule (1975) defined children with achievement problems according to whether they were "backward readers," who read at IQ-appropriate levels, or "specific reading disabled," who read below expected levels according to their IQ scores. To date, however, relatively little empirical evidence exists showing that similarly defined children differ on measures other than IQ (Morris & Satz, 1986). Shaywitz, Shaywitz, Barnes, and Fletcher (1986) compared the influence of various definitions on the selection of children as learning disabled in an epidemiological study. Although variations in the use of IQ indices and definitions resulted in different children being identified as learning disabled, few differences in cognitive ability were apparent among children grouped as learning disabled according to various definitions. There were also few differences among children defined as learning disabled whose scores were discrepant or not discrepant with IQ.

In order to develop more adequate, empirically based definitions of reading disability, systematic analyses of the characteristics of various definitional criteria leading to the identification of children with reading disabilities must be completed. There is persistent interest in typologies of reading disabilities and other classification hypotheses (Rourke, 1985). However, research addressing the criteria for defining a disorder is a prerequisite to any search for subtypes, since the emergence of a typology will depend in part on how children are entered or excluded from the sample (Morris & Fletcher, 1988). Different subtypes may emerge based on issues as simple as the IQ cutoff and achievement cutoff scores defining

the sample.

The present study expands upon Shaywitz et al. (1986) by employing a large, clinic-based (but school-referred) sample in an attempt to study (a) the effects of variations in definitions and (b) whether differences in neuropsychological skills emerge between children whose reading skills are discrepant or not discrepant with IQ.

METHOD

Subjects

The children for this study were obtained from a data base of over 2,500 cases representing children referred for evaluation of learning disability in Windsor, Ontario. Each child received a comprehensive neuropsychological evaluation (Rourke, 1981; Rourke, Fisk, & Strang, 1986), and the Wechsler Intelligence Scale for Children (WISC) (Wechsler, 1949) and the Wide Range Achievement Test (WRAT) (Jastak & Jastak, 1965) were administered. For this study, children were selected ranging in age from 9 years to 14 years, with Full Scale WISC IQ scores above 79. These children were free of sensory, acquired neurological, and other problems traditionally used as exclusionary criteria, resulting in a total sample of 1,069 children. The sample, 74% of whom were male, were predominantly white, middle class children who averaged 11 years, 4 months in age, with a mean WRAT reading standard score of 89.3 ($SD = 14.8$) and mean WISC FSIQ of 98.5 ($SD = 10.5$).

Procedures

Two different definitions were used to identify children as reading disabled, based on the word recognition score from the WRAT and the WISC FSIQ. The first definition employed a cutting-score approach that did not correct for regression artifact. Children were defined as reading disabled if their FSIQ exceeded 79 and their WRAT reading standard score was below 93. In addition, children were categorized according to whether reading scores were consistent with or inconsistent with FSIQ using a criterion of 15 points. A child was considered "dis-

crepant" if his or her WRAT score was 15 points or lower than FSIQ. This definition corresponds directly with criteria used by the Texas Education Agency (TEA) and other states to define eligibility for reading disabilities services. Liberal criteria in terms of relative severity of word recognition deficit and IQ were used in the hope of capturing the largest possible sample unbiased according to selection variables.

RESULTS

Table 1 presents sample sizes for the resultant 2 x 2 matrix of children who were classified according to definition and IQ-reading discrepancies. Note that there is a small group of children ($n = 36$) with reading standard scores greater than 92 whose reading score was at least 15 points below FSIQ. The other children are distributed fairly evenly across the matrix. About 34% had reading standard scores below 93 that were at least 15 points below their FSIQ scores, with 30% regarded as not impaired in reading. Some 32% of the children exhibited poor reading and low average FSIQ, but would not qualify for services.

The second definition corrected for regression artifact inherent in the comparison of WRAT and WISC scores using a formula adopted by Shaywitz et al. (1986). This formula is a variation on those provided by Reynolds (1984) and corrects for the correlation of FSIQ and WRAT Reading in the specific population of interest. This formula also takes into account the nonuniform variability in predicted reading scores at different values of IQ. Using a comparable definition of discrepancy, Table 2 presents the resultant 2 x 2 matrix. It is apparent that the distribution of children across the four categories is different from that in Table 1. More children are identified as discrepant and fewer as nondiscrepant. In terms of overlap between the two definitions, 70 children (7%) with reading standard scores below 93 become eligible for services using the regression-based definition who were not eligible under the cutoff score definition. These children were generally lower on FSIQ ($M = 87.7$; $SD = 4.5$) and WRAT reading ($M = 75.7$; $SD = 3.4$) than were

the group of discrepant readers identified in the first analysis. However, 22 (2%) children who had reading standard scores below 93 and 15-point discrepancies between IQ and reading were no longer eligible under the regression-based definition. These children had a mean WISC FSIQ of 106.2 ($SD = 3.0$) and mean WRAT reading standard score of 89.7 ($SD = 3.0$). Of the 36 higher-IQ-but-discrepant achievement children eligible under the cutoff score criterion, only 5 remain eligible when the definition takes into account regression effects. Thus, the regression-based criteria make 70 "poor readers" eligible and eliminate eligibility for 22 children with word recognition scores below 93 but within the range expected given their IQ, and for 31 children with age-appropriate word recognition scores ($M = 97.1$; $SD = 3.6$) and higher FSIQ ($M = 114.6$; $SD = 4.2$).

Effects of Variations in Definition

A second set of analyses addressed whether variations in definition produced differences in ability structure among groups formed with different definitional criteria. To address this possibility, low achieving children who were discrepant and not discrepant under the two definitions were compared on a set of tests derived from a modification of the Halstead-Reitan Neuropsychological Battery for Children (HRB) (Rourke et al., 1986). These measures constitute a representative sample of neuropsychological skills and abilities that are ordinarily administered in a comprehensive evaluation of children with learning disabilities (Rourke, 1981). The linguistic and auditory-perceptual measures are especially sensitive to the reliable discrimination of children with reading disability from nondisabled children and from children with other types of learning disabilities (Rourke, 1978, 1981; Rourke et al., 1986).

Ten tests from the HRB were used. These tests are presented in Table 3 along with a summary of the constructs measured by each task. These constructs were defined according to recent maximum-likelihood factor analyses of the test battery in this sample completed by our group. It is apparent that these tests measure a variety of abilities frequently

TABLE 1
Distribution, Mean IQ, and Word Recognition Scores of Children Categorized According to Reading Standard Scores and Raw Discrepancies

	Reading Standard Score	
	≤ 92	> 92
Discrepant	$n = 360$ (34%)	$n = 36$ (4%)
	FSIQ = 101.8 (8.7)	FSIQ = 114.9 (6.3)
	RdSS = 78.4 (8.0)	RdSS = 96.1 (5.1)
Not Discrepant	$n = 347$ (32%)	$n = 326$ (30%)
	FSIQ = 90.6 (6.0)	FSIQ = 101.3 (10.8)
	RdSS = 83.9 (5.7)	RdSS = 106.4 (12.2)

Note. FSIQ = Full Scale IQ on Wechsler Intelligence Scale for Children; RdSS = reading standard score on Wide Range Achievement Test.

TABLE 2
Distribution, Mean IQ, and Word Recognition Scores of Children Categorized According to Reading Standard Scores and Regression-Based Discrepancies

	Reading Standard Score	
	≤ 92	> 92
Discrepant	$n = 408$ (38%)	$n = 5$ (1%)
	FSIQ = 99.2 (9.8)	FSIQ = 99.2 (9.8)
	RdSS = 77.3 (7.2)	RdSS = 90.2 (9.1)
Not Discrepant	$n = 299$ (28%)	$n = 357$ (33%)
	FSIQ = 92.4 (7.1)	FSIQ = 102.5 (11.0)
	RdSS = 86.2 (4.2)	RdSS = 105.6 (12.0)

Note. FSIQ = Full Scale IQ on Wechsler Intelligence Scale for Children; RdSS = reading standard score on Wide Range Achievement Test.

TABLE 3
Modified Halstead-Reitan Neuropsychological Tests by Factor Structure

Test	Factor
1. Category Test	Executive Functions, Spatial Relations
2. Speech-Sounds Perception Test	General Language Acoustic Language
3. Auditory Closure Test	General Language, Acoustic Language
4. Sentence Memory Test	General Language, Acoustic Language
5. Verbal Fluency Test	General Language, Acoustic Language
6. Finger Tapping Test	Simple Motor
7. Grooved Pegboard Test	Eye-Hand Coordination, Spatial Relations
8. Tactual Performance Test	Spatial Relations, Executive Function, Eye-Hand
9. Trail Making Test, Parts A and B	Executive Function
10. Target Test	Spatial Relations, Eye-Hand

impaired in children with reading disabilities, including language, perceptual, and motor skills.

All scores were converted to age standardized scores using available normative data. For each comparison of discrepant and nondiscrepant readers, the data were subjected to a linear discriminant function analysis with follow-up interpretation of the discriminant functions to determine variable contributions to group separation. Identification rates (i.e., proportion of children correctly placed into reading groups based on HRB tests) were computed as an estimate of effect size.

Cutoff Versus Regression-Based Comparisons

Table 4 summarizes the results of the comparison of discrepant and nondiscrepant poor readers on the 10 HRB tasks using the two definitions. A significant discriminant function was obtained for each comparison: cutoff, $F(10,696) = 20.1$, $p < .001$; regression, $F(10,696) = 15.3$, $p < .0001$. In terms of variable selection, it is apparent that results are similar for both comparisons, with considerable overlap in those variables maximally separating the groups. In both comparisons, the Speech-Sounds Perception Test, a measure of acoustic language that involves recognition of the visual representations of aurally presented nonsense words, is the best discriminator. The basic difference in the two comparisons is the emergence of the Category Test (executive functions, spatial relations factors) as an important discriminator in the regression-based comparison, and the Target Test (spatial relations, eye-hand coordination factors) in the cutoff score comparison. For the cutoff score comparison, the nondiscrepant group was lower on all variables except for the Sentence Memory Test (rote language factor). In the regression-based comparison, the distinctions seemed more meaningful, with the discrepant group scoring lower only on the Speech-Sounds Perception Test (acoustic language factor) and the nondiscrepant group scoring lower on a variety of tasks involving concept formation and spatial relations. Identification rates of .69 in both analyses show that effect sizes are small but comparable.

TABLE 4
Comparisons of Standardized Discriminant Function Weights Maximally Separating Discrepant and Nondiscrepant Children with Reading Disabilities on Neuropsychological Tests

Variable	Definition			
	Cutoff Score		Regression-based	
	Unadjusted	Adjusted ^a	Unadjusted	Adjusted ^a
Category Test	.35	.07	.27	.05
Speech Sounds Perception Test	-.59	-.14	.78	-.14
Auditory Closure Test	.10	-.11	.29	-.19
Sentence Memory Test	.47	.29	-.39	.34
Verbal Fluency Test	-.51	.86	-.03	.87
Finger Tapping Test	.05	.20	.09	.19
Grooved Pegboard Test	.36	.16	.34	.10
Tactual Performance Test	.42	.23	.31	.16
Trail Making Test, Parts A and B	.10	.06	-.06	.08
Target Test	-.29	-.05	.33	-.15

^aCovaried for WISC Full Scale IQ and age.

Influence of IQ, Age, and Achievement

The 10 HRB tests covary with IQ and age. Although all scores were age-corrected, we chose to covary age to eliminate any possible influence due to non-specific sample characteristics that might have affected the applicability of the age norms first used to adjust scores. It was not apparent that covariation for age had significant influence on the results, so these results are not separately reported. Covariation for IQ was conducted to estimate the magnitude and nature of group differences when variability due to FSIQ scores was eliminated. The results of these discriminant function analyses are also summarized in Table 4. The adjusted discriminant functions were significant: cutoff, $F(10,694) = 9.5$, $p < .0001$; regression, $F(10,694) = 14.3$, $p < .0001$. Three variables contributed to group separation in both analyses: Verbal Fluency, Sentence Memory, and Grooved Pegboard. The effect size (.65) is comparable to the unadjusted analyses (.69), showing that removal of IQ had little effect on the magnitude of group differences. The group profiles are similar in both analyses, with the discrepant group scoring lower on the Verbal Fluency Test

(acoustic language factor) and the nondiscrepant group scoring lower on the Grooved Pegboard Test (eye-hand, spatial relations factors). These results show that removing IQ differences leads to a common set of discriminators between the two definitions, with no difference in effect size. Differences in variable weights across analyses generally reflect relationships of various measures with FSIQ, not group membership.

DISCUSSION

These analyses addressed two questions. The first question was the influence of different definitions of reading disability in terms of which children were identified as "reading disabled." The second question concerned differences on external variables between children with reading problems whose scores were discrepant or not discrepant with IQ-based expectations.

For the first question, these results show that variations in definitional criteria lead to differences in which children are designated as reading disabled. With a cutoff score approach, IQ score differences between discrepant and nondiscrepant children are larger, word recogni-

tion score differences are smaller, and many lower IQ children are not identified as eligible for services. In contrast, more children become eligible for services when regression-based procedures are used, with some higher IQ, better reading children becoming ineligible. The IQ differences are negligible, but differences in word recognition scores are larger. Hence, one conclusion may be that use of unadjusted IQ-achievement discrepancies is biased against children who score lower on IQ tests.

In terms of the second question, differences between discrepant and nondiscrepant children on external variables are significant for both definitions, but are small, difficult to characterize, and tend to dissipate when IQ and age are used as covariates. The main influence of using a regression-based procedure is to more clearly polarize group differences on external measures (despite reducing IQ differences) by adding lower IQ children. The discrepant group displays lower performance than the nondiscrepant group only on measures of acoustic language. As such, the results are similar to those of Jorm, Share, MacLean, and Matthews (1986), who found that specific problems with phonological language discriminated a small group of poor readers with discrepancies in IQ and reading from a group of nondiscrepant poor readers.

It is possible that the regression-based definitions produce a more homogeneous group of poor readers. However, further research using measures more clearly related to reading proficiency may be necessary to elucidate these differences. Individual HRB tests are factorially confounded, measuring several different abilities. Some of these abilities may be related to reading disability and other abilities to some other less specific (but adaptive) characteristics of the child. Whether group differences between discrepant and nondiscrepant children would be clearer with other measures is an empirical question. However, given the small effect size and the heterogeneity of the groups, it is not likely that any other set of measures would produce clearer differences between the groups, particularly given the absence of any theoretical reason to expect group differences.

Eliminating variability due to age and IQ did not significantly alter effect size,

but did alter the number and nature of discriminating variables. Since the scores were age-corrected, IQ was the principal covariate. The neuropsychological variables discriminated almost as robustly when IQ was eliminated from overall variability, showing that such measures capture variability not reflected in IQ tests (Taylor et al., 1984). However, the nature of group differences is quite similar between the two definitions. Additional research should address more precisely the role and contribution of IQ scores to definitions of reading disability. The fact that discrepant and nondiscrepant reading groups show very small differences when IQ is covaried implies that the IQ differences may not be important.

The results of this study demonstrate that the specificity of a reading disability depends in part on how it is defined. Sample constituents change when definitional criteria are varied. However, the notion that intelligence scores are determinants of reading proficiency or indices of potential is not supported by this study. IQ scores are actually lower when regression-based procedures are used, but the groups seem to become more homogeneous. The next step in this type of research would be systematic variation of other assumptions inherent in definitions of learning disabilities. For example, Shaywitz et al. (1986) explored various IQ cutoffs and the influence of varying levels of severity of reading (i.e., achievement cutoffs). Degree of discrepancy (1 *SD*, 1.5 *SD*, 2 *SD*) could be varied. In each instance, a set of external variables should be available to provide criteria for comparisons. Treatment outcomes would be especially relevant.

This type of classification research is straightforward, time-consuming, and perhaps to some, dry and boring. However, classification research of this type is fundamental to any scientific endeavor and essential to future progress in studying and remediating children with reading disabilities (Fletcher & Morris, 1986). Determining the influence of variations in definitional criteria is a prerequisite for adequate understanding of the cognitive and biological correlates of reading disabilities. Statements about specificity and definition remain empirical questions in need of considerable research.

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AUTHORS' NOTES

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