11-1-2006

Plant Reproductive Allocation Predicts Herbivore Dynamics across Spatial and Temporal Scales

Tom E. X. Miller  
*University of Nebraska - Lincoln*

Andrew J. Tyre  
*University of Nebraska - Lincoln*, atyre2@unl.edu

Svata M. Louda  
*University of Nebraska - Lincoln*, slouda1@unl.edu

Follow this and additional works at: [http://digitalcommons.unl.edu/bioscifacpub](http://digitalcommons.unl.edu/bioscifacpub)

Part of the [Life Sciences Commons](http://digitalcommons.unl.edu/bioscifacpub)

[http://digitalcommons.unl.edu/bioscifacpub/55](http://digitalcommons.unl.edu/bioscifacpub/55)

This Article is brought to you for free and open access by the Papers in the Biological Sciences at DigitalCommons@University of Nebraska - Lincoln. It has been accepted for inclusion in Faculty Publications in the Biological Sciences by an authorized administrator of DigitalCommons@University of Nebraska - Lincoln.
Plant Reproductive Allocation Predicts Herbivore Dynamics across Spatial and Temporal Scales

Tom E. X. Miller, Andrew J. Tyre, and Svata M. Louda

1. School of Biological Sciences, University of Nebraska, Lincoln, Nebraska 68588; 2. School of Natural Resources, University of Nebraska, Lincoln, Nebraska 68583

Submitted December 2, 2005; Accepted June 14, 2006; Electronically published October 5, 2006

Abstract: Life-history theory suggests that iteroparous plants should be flexible in their allocation of resources toward growth and reproduction. Such plasticity could have consequences for herbivores that prefer or specialize on vegetative versus reproductive structures. To test this prediction, we studied the response of the cactus bug (Narnia pallidicornis) to meristem allocation by tree cholla cactus (Opuntia imbricata). We evaluated the explanatory power of demographic models that incorporated variation in cactus relative reproductive effort (RRE; the proportion of meristems allocated toward reproduction). Field data provided strong support for a single model that defined herbivore fecundity as a time-varying, increasing function of host RRE. High-RRE plants were predicted to support larger insect populations, and this effect was strongest late in the season. Independent field data provided strong support for these qualitative predictions and suggested that plant allocation effects extend across temporal and spatial scales. Specifically, late-season insect abundance was positively associated with interannual changes in cactus RRE over 3 years. Spatial variation in insect abundance was correlated with variation in RRE among five cactus populations across New Mexico. We conclude that plant allocation can be a critical component of resource quality for insect herbivores and, thus, an important mechanism underlying variation in herbivore abundance across time and space.

Keywords: demographic model, herbivory, insect-plant interactions, Opuntia, population dynamics, resource allocation.

Life-history functions, such as plant growth and reproduction, typically incur physiological and demographic costs (Bazzaz and Grace 1997). Consequent trade-offs between functions are axiomatic to plant life-history theory (Harper 1977; Silvertown and Dodd 1999). Such theory predicts that iteroparous perennials—long-lived plants capable of repeated reproductive bouts—should be flexible in their relative allocation of resources to growth versus reproduction (Bonser and Aarssen 1996; Bazzaz and Grace 1997).

Resource allocation decisions can be quantified in the currency of meristems, the primordial tissues from which vegetative and reproductive organs arise (Geber 1990; Bonser and Aarssen 1996; Geber et al. 1997; Olejniczak 2001; Hartemink et al. 2004). The total number of active meristems available at a given time is finite, and once committed to a particular function, a meristem cannot be redirected to another function (Watson 1984). For non-clonal, iteroparous perennial plants, this can present an acute strategic dilemma. Allocation of meristems to current reproduction has the potential for immediate fitness benefits, but it constrains growth and, consequently, future reproductive potential. Conversely, the allocation of meristems to production of vegetative structures limits current reproductive output, but long-term gains in size, survival, and future fecundity could outweigh current losses (Geber 1990; Hartemink et al. 2004).

Meristem allocation strategies to cope with these trade-offs can vary widely among species, among populations within species, among individuals within populations, and within individuals across time; such variation is consistent with theory and may be driven by genetic and/or environmental factors (Geber 1990; Bowers 1996; Duffy et al. 1999; Reekie 1999; Hartemink et al. 2004; Mendez and Karlsson 2004). These patterns lead to the prediction that variation in the production of vegetative versus reproductive parts should have consequences for herbivores that prefer or specialize on specific plant structures. Yet, plasticity in plant resource allocation has rarely been examined as a potential driver of variation in the abundance of higher...
The coreid bug \textit{Narnia pallidicornis} specializes on \textit{Opuntia} cacti (Mann 1969). There are two generations within a year, and juveniles and adults co-occur on cacti throughout the growing season. Adults overwinter among debris near the base of the plant, and females deposit eggs on cactus spines from March to May. The first cohort of nymphs reaches maturity in May–June, and the second, larger, cohort matures in August–September (Mann 1969). Juveniles are flightless and complete their development on a single host plant (Mann 1969). Early-season abundance on a plant is not correlated with late-season abundance from the previous year (data from 2003 to 2004: \( n = 47, r = 0.11, P = A7 \)), suggesting that \textit{Narnia} dynamics at the single plant level are independent across adjacent years.

The tree cholla (~\textit{Opuntia imbricata}) is a perennial arborescent cactus, native to New Mexico, west Texas, and southern Colorado (Kinraide 1978; Benson 1982), that produces cylindrical stem segments. Early in the growing season (mid-May to early June in central New Mexico), meristems appear in clusters at the terminal ends of stem segments produced the previous year, and these develop into either reproductive or vegetative structures. Merists allocated to reproduction develop into flower buds in late spring; the fruits ripen and fall off of plants in early fall. Meristems allocated to new stem segments grow from May to August and produce new meristems the following spring.

We quantified plant allocation as the proportion of total available meristems that developed into reproductive structures, which we call relative reproductive effort (RRE). This metric is bound by 0 and 1, where a plant with RRE = 0 allocates all meristems to growth, a plant with RRE = 1 allocates all meristems to reproduction, and a plant with RRE = 0.5 allocates equally between growth and reproduction. We focus on RRE rather than, for example, absolute numbers of reproductive structures for two important reasons. First, any variation in flower production among similarly sized plants ultimately must be driven by variable individual allocation strategies; our aim was to quantify these strategies and their consequences directly rather than their proxies or corollaries. Second, trade-offs between life-history functions assume a central role in current thinking about plant resource allocation (Bazzaz and Grace 1997; Obeso 2002). Because RRE explicitly accounts for such trade-offs, we think that this approach allows our results to be interpreted in light of, and integrated into, existing theory.

\section*{Methods}

\subsection*{Natural History of the Study Organisms}

The coreid bug \textit{Narnia pallidicornis} specializes on \textit{Opuntia} cacti (Mann 1969). There are two generations within a year, and juveniles and adults co-occur on cacti throughout the growing season. Adults overwinter among debris near the base of the plant, and females deposit eggs on cactus spines from March to May. The first cohort of nymphs reaches maturity in May–June, and the second, larger, cohort matures in August–September (Mann 1969). Juveniles are flightless and complete their development on a single host plant (Mann 1969). Early-season abundance on a plant is not correlated with late-season abundance from the previous year (data from 2003 to 2004: \( n = 47, r = 0.11, P = A7 \)), suggesting that \textit{Narnia} dynamics at the single plant level are independent across adjacent years.

The tree cholla (~\textit{Opuntia imbricata}) is a perennial arborescent cactus, native to New Mexico, west Texas, and southern Colorado (Kinraide 1978; Benson 1982), that produces cylindrical stem segments. Early in the growing season (mid-May to early June in central New Mexico), meristems appear in clusters at the terminal ends of stem segments produced the previous year, and these develop into either reproductive or vegetative structures. Merists allocated to reproduction develop into flower buds in late spring; the fruits ripen and fall off of plants in early fall. Meristems allocated to new stem segments grow from May to August and produce new meristems the following spring.

We quantified plant allocation as the proportion of total available meristems that developed into reproductive structures, which we call relative reproductive effort (RRE). This metric is bound by 0 and 1, where a plant with RRE = 0 allocates all meristems to growth, a plant with RRE = 1 allocates all meristems to reproduction, and a plant with RRE = 0.5 allocates equally between growth and reproduction. We focus on RRE rather than, for example, absolute numbers of reproductive structures for two important reasons. First, any variation in flower production among similarly sized plants ultimately must be driven by variable individual allocation strategies; our aim was to quantify these strategies and their consequences directly rather than their proxies or corollaries. Second, trade-offs between life-history functions assume a central role in current thinking about plant resource allocation (Bazzaz and Grace 1997; Obeso 2002). Because RRE explicitly accounts for such trade-offs, we think that this approach allows our results to be interpreted in light of, and integrated into, existing theory.

\section*{Stage-Structured Insect Dynamics and Host Plant Allocation Effects}

We describe the per-plant dynamics of juvenile and adult cactus bugs with the following model structure:
\[ I_{t+\Delta t} = s_j(1 - T)I_j + f(s_jA_j + c), \]
\[ A_{t+\Delta t} = s_jT I_j + s_c A_c + c, \]  
\[ \text{(1)} \]

where \( I_j \) and \( A_j \) are juvenile and adult bug abundances, respectively, at time \( t \); \( s_j \) and \( s_c \) are the survival probabilities of juveniles and adults, respectively; \( T \) is the probability of transition from the juvenile to the adult stage; \( f \) is adult realized fecundity (juveniles per adult); and \( c \) is the adult colonization rate (adults per \( \Delta t \)). There are multiple time steps within a growing season (i.e., \( \Delta t < 1 \) year), and this model structure assumes that juveniles cannot transition and reproduce within a single time step. Because juvenile bugs are flightless and complete development on single plants, there is no juvenile movement in the model. Because detecting eggs and differentiating among instars are difficult and time consuming in the field, the egg stage is not explicitly included in the model, and all immature stages are considered collectively in the juvenile class. The colonization parameter is density independent in order to account for midseason appearance of insects on previously unoccupied plants. Adult emigration also occurs (clearly, colonists are coming from somewhere), but it is implicitly incorporated into the model because emigration and mortality could not be differentiated in time series data. Thus \( s_c \) represents the joint probability of an adult surviving and remaining on a host plant to be detected at the next time step.

Using this basic structure, we examined the relative explanatory power of 16 models that incorporated realistic variation in the five demographic parameters (table 1). In these models, each parameter could take one of four forms. These were (1) constant across time and host plants, (2) unique to each time step (e.g., colonization: \( c = c_t \)), (3) unique to each host plant as a constant linear function of its RRE (e.g., \( c = c_x \times \text{RRE} \)), and (4) unique to each host plant as a time-varying linear function of cactus RRE (e.g., \( c = c_x \times \text{RRE} \)). Because resource quality changes over the season and because the probability of certain demographic events may vary with population age structure, we expected that the time-varying models would provide a better fit to field data.

Each of our candidate models (table 1) allows a single parameter to vary with time or with host allocation (as opposed to multiple parameters in combination). This approach allowed us to evaluate a broad range of demographic hypotheses while keeping the total number of models and the number of parameters in each relatively small. In this way, we focus on the issues of whether RRE influences insect dynamics and, if so, which demographic transition best explains an RRE effect rather than trying to differentiate the specific mechanics of such an effect.

\textbf{Data Collection and Model Fitting}

Data used to test the relative explanatory power of the competing models were collected during spring and summer 2004 in a Chihuahuan Desert grassland at the Sevilleta National Wildlife Refuge, a long-term ecological research (LTER) site in central New Mexico (34°20'5.3"N, 106°37'53.2"W). Annual precipitation at the study site averages 245 mm/year, most of which falls from July to September (Sevilleta LTER Program meteorological data, 1989–1999). In early May 2004, we located and tagged 80 mature tree cholla at the Sevilleta. Numbers of juvenile and adult 	extit{Narna} on each plant were counted on five sampling dates 25–30 days apart: May 13–15, June 7–9, July 2–4, July 24–26, and August 16–18, 2004. Any eggs observed were not included in the juvenile counts. For each plant, we also counted the numbers of newly produced stem segments and flower buds in June and used these counts to calculate cactus RRE. These plants showed a wide range of meristem allocation strategies (minimum RRE = 0, maximum RRE = 0.86). Using sequential counts of stem segments from the most recent growth to the root crown, we estimated that these plants were 10.3 ± 0.21 (mean ± SE) years old at the time sampled. This method assumed no loss of segments and therefore provided a minimum age estimate.

We used maximum likelihood methods to confront the 16 candidate models with data (Hilborn and Mangel 1997) and Akaike’s Information Criterion (AIC) to select the one with the most explanatory power (Burnham and Anderson 1998). We defined the model time step based on the schedule of data collection; that is, \( \Delta t \approx 25 \) days. This interval is greater than the duration of the egg stage (egg duration of a closely related cactus-feeding coreid is 12 days; Mann 1969), and so our estimate of fecundity (juveniles per adult) does not require a time lag in the model.

We first generated 16 season-wide projections of 	extit{Narna} abundance on each plant, each corresponding to one of the models in table 1, using the May sampling date as \( t_0 \). Projections at each subsequent time step were based on observed rather than predicted abundances; thus, the model included process error but not observation error (Hilborn and Mangel 1997). We considered the presence of an insect on a plant as a random event and used a natural log–transformed Poisson function, which gives the likelihood of each observation (\( x \)), given the value predicted by the model (\( \mu \)). We calculated the total negative log likelihood (\( L \)) of each model by summing the likelihoods over all sampling dates (\( t = 1, \ldots, 4 \)) and plants (\( m = 1, \ldots, 80 \)):

\[ L = \sum_{m=1}^{80} \sum_{t=1}^{4} \mu_m - x_m \ln \mu_m + \ln (x_m!). \]  
\[ \text{(2)} \]
We minimized $L$ over all model parameters using the Solver feature in Microsoft Excel. The survival and transition probabilities were bound by 0 and 1, and colonization rate and fecundity were constrained to be $\geq 0$ using logistic and exponential transformations, respectively. The constraint on colonization ensured positive predicted values of insect abundance ($L$ is undefined for $\mu \leq 0$).

For each model, we calculated AIC, a parsimony-based statistic that imposes an increasing penalty on $L$ as the number of parameters increases. We then calculated AICc values and Akaike weights for each model according to equations given by Burnham and Anderson (1998). The AICc values are corrected for sample size ($n = 80$ plants), and Akaike weights can be interpreted as the proportion of evidence in favor of each model within a given set (Burnham and Anderson 1998). We did not use model averaging to make our predictions because the Akaike weight for the best model was substantively greater than all other models combined.

**Data Collected to Evaluate Model Predictions**

To evaluate qualitative model predictions for variation in insect abundance in relation to host plant resource allocation, we collected three additional, independent data sets on *Narnia* abundance and *O. imbricata* RRE across years and both local and regional spatial scales. First, we monitored an additional 117 *O. imbricata* cacti at the Sevilleta during summer 2004. These data were used to evaluate predictions among plants for the same site and year as the model fitting data. Second, a subset of these plants ($n = 47$) was also visited and measured twice during the previous year (2003) and twice the following year (2005). We used these observations to determine whether inter-annual variation in cactus bug abundance was associated with variation in tree cholla resource allocation. Because our model predicted within-season variation in host allocation effects (see “Results”), if an association occurred, we also asked whether it differed between early- and late-season cohorts. Third, in 2004 we examined five *O. imbricata* populations along a south-north transect in central New Mexico (see map in fig. 2B) to determine whether large-scale spatial variation in *Narnia* abundance was associated with among-population variation in host plant RRE. Further information on data collection methods and statistical analyses is given in appendix A in the online edition of the *American Naturalist*.

**Results**

We used field data on cactus RRE and insect abundance to evaluate the explanatory power of 16 models of cactus

---

**Table 1: Candidate models of cactus bug (*Narnia pallidicornis*), population dynamics on tree cholla (*Opuntia imbricata*), and the results of model fitting to a subset of the field data**

<table>
<thead>
<tr>
<th>Model</th>
<th>Adult survival</th>
<th>Juvenile survival</th>
<th>Colonization</th>
<th>Fecundity</th>
<th>Transition</th>
<th>No. parameters</th>
<th>$\Delta$AICc</th>
<th>$w_i$</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>$s_a$</td>
<td></td>
<td>$c$</td>
<td>$f$</td>
<td>$T$</td>
<td>5</td>
<td>586.96</td>
<td>3.5E-128</td>
</tr>
<tr>
<td>2</td>
<td>$s_a$</td>
<td></td>
<td>$c$</td>
<td>$f$</td>
<td>$T$</td>
<td>8</td>
<td>489.09</td>
<td>6.3E-107</td>
</tr>
<tr>
<td>3</td>
<td>$s_a$ (RRE)</td>
<td></td>
<td>$c$</td>
<td>$f$</td>
<td>$T$</td>
<td>5</td>
<td>460.94</td>
<td>8.1E-101</td>
</tr>
<tr>
<td>4</td>
<td>$s_a$ (RRE)</td>
<td></td>
<td>$c$</td>
<td>$f$</td>
<td>$T$</td>
<td>8</td>
<td>394.84</td>
<td>1.8E-86</td>
</tr>
<tr>
<td>5</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f$</td>
<td>$T$</td>
<td>8</td>
<td>311.68</td>
<td>2.1E-68</td>
</tr>
<tr>
<td>6</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f$</td>
<td>$T$</td>
<td>5</td>
<td>526.18</td>
<td>5.5E-115</td>
</tr>
<tr>
<td>7</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f$</td>
<td>$T$</td>
<td>8</td>
<td>414.20</td>
<td>1.1E-90</td>
</tr>
<tr>
<td>8</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f$</td>
<td>$T$</td>
<td>8</td>
<td>544.95</td>
<td>4.6E-119</td>
</tr>
<tr>
<td>9</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$ (RRE)</td>
<td>$f$</td>
<td>$T$</td>
<td>5</td>
<td>369.97</td>
<td>4.6E-81</td>
</tr>
<tr>
<td>10</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$ (RRE)</td>
<td>$f$</td>
<td>$T$</td>
<td>8</td>
<td>284.21</td>
<td>1.9E-62</td>
</tr>
<tr>
<td>11</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f_T$</td>
<td>$T$</td>
<td>8</td>
<td>240.25</td>
<td>6.8E-53</td>
</tr>
<tr>
<td>12</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f_T$ (RRE)</td>
<td>$T$</td>
<td>5</td>
<td>416.87</td>
<td>3.0E-91</td>
</tr>
<tr>
<td>13</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f_T$ (RRE)</td>
<td>$T$</td>
<td>8</td>
<td>0.00</td>
<td>.99</td>
</tr>
<tr>
<td>14</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f$</td>
<td>$T_T$</td>
<td>8</td>
<td>466.64</td>
<td>4.7E-102</td>
</tr>
<tr>
<td>15</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f_T$ (RRE)</td>
<td>$T$</td>
<td>5</td>
<td>582.99</td>
<td>2.5E-127</td>
</tr>
<tr>
<td>16</td>
<td>$s_a$</td>
<td>$s_b$</td>
<td>$c$</td>
<td>$f_T$ (RRE)</td>
<td>$T$</td>
<td>8</td>
<td>441.65</td>
<td>1.3E-96</td>
</tr>
</tbody>
</table>

Note: In these models, each parameter was constant, dependent on time, or dependent on host plant relative reproductive effort (RRE) as either a constant or a time-varying linear function. General model structure is shown in the text (eq. [1]). The table shows the functional form of the demographic parameters; the number of parameters in each model; the $\Delta$AICc value, which gives the difference between the Akaike’s Information Criterion value adjusted for sample size (AICc) of each model and the lowest AICc of all models; and Akaike weights ($w_i$), which give the proportional weight of evidence in favor of each model. We did not include intercepts in the models incorporating RRE. Models with such extra parameters either did not converge or led to biologically implausible values with large confidence limits for all parameters. Model 13, with host $\times$ time-dependent fecundity (shown in bold), is the most likely, given the data. See table 2 for maximum likelihood parameter estimates.
Discussion

Variation in meristem allocation is widespread and critically important in long-term plant performance, yet plant allocation patterns have seldom been evaluated for their effects on insect herbivore abundance or dynamics. Here, we found that relative allocation of meristems to reproduction versus to vegetative growth had strong predictive ability across temporal and spatial scales for the cactus bug on tree cholla cacti. These results suggest that plant resource allocation is an important aspect of host plant quality for *Narnia* and that plant allocation can influence and help explain variation in herbivore abundance within and among plant populations.

The model of cactus bug population dynamics that best described our field data (table 1, model 13) predicted that greater allocation of meristems toward reproductive structures led to greater per-plant bug abundances and that this effect intensified from early to late in the growing season (fig. 1). The model identified fecundity (juveniles per adult) as the insect population parameter most responsive to plant allocation. There are at least two possible explanations for the temporal variation in parameter estimates that we observe (table 2). First, insect fecundity may be related to plant resource quality, which likely increases throughout the season as meristems differentiate (insect population growth begins before complete differentiation) and fruits ripen. Second, population age structure (within stage structure) could cause the time-dependent model to provide the best fit, independent of resource seasonality. For example, adults observed in the first census period had reached maturity the previous fall, whereas those observed at the second census had likely just matured from the first cohort of juveniles. Thus, any differences in fecundity between old and young adults could also drive temporal variation in parameter estimates. Our data do not allow us to distinguish between these possibilities, and the temporal variation in model structure is difficult to interpret with certainty.

Additional data, independent of those used for model fitting, supported the main model predictions of the importance of plant allocation. Increases in tree cholla RRE over 3 years coincided with significant increases in *Narnia* abundance that were greater for the late- than early-season cohorts (fig. 2A), as predicted. Spatially, tree cholla populations that, on average, allocated proportionally more meristems toward reproduction supported larger cactus bug populations (fig. 2B), as predicted. It is important to

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Point estimate (95% confidence limits)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Juvenile survival (<em>s</em>)</td>
<td>0.31 (.24, .38)</td>
</tr>
<tr>
<td>Adult survival (<em>s</em>)</td>
<td>0.28 (.24, .34)</td>
</tr>
<tr>
<td>Transition (<em>T</em>)</td>
<td>0.16 (.08, .25)</td>
</tr>
<tr>
<td>Colonization (<em>c</em>)</td>
<td>0.86 (.76, .97)</td>
</tr>
<tr>
<td>Fecundity (<em>f</em>)</td>
<td></td>
</tr>
<tr>
<td>May</td>
<td>10.22 (8.6, 12.14) × RRE</td>
</tr>
<tr>
<td>June</td>
<td>1.31 (.84, 1.93) × RRE</td>
</tr>
<tr>
<td>July</td>
<td>9.04 (7.76, 10.56) × RRE</td>
</tr>
<tr>
<td>August</td>
<td>18.39 (15.92, 21.29) × RRE</td>
</tr>
</tbody>
</table>

Note: RRE = relative reproductive effort.
Note that while the independent data support the qualitative predictions of the model, they provide no information about underlying mechanisms. Thus, we do not know whether other demographic parameters (e.g., migration, survival) or other functional forms of fecundity drive these patterns across years and sites. In fact, neither the cross-year nor the cross-site insect-RRE relationship approximates a 0 intercept (as our single-year site model does), suggesting that other processes may operate at larger scales.

While multiple factors will contribute to variation in the abundance of specialist insect herbivores across time and space, our analyses strongly suggest that host plant resource allocation pattern is important. We conducted identical model fitting analyses in which RRE was replaced with other plant measures (age, size, number of reproductive structures) and found that RRE had the most predictive ability (see app. B in the online edition of the American Naturalist). We also examined differences in winter climate to explain interannual variation. However, differences in climate are unlikely to directly account for the increase in *Narnia* abundance during this study because sizes of the first cohort were similar across years (fig. 2A), suggesting that rates of overwinter survival were also similar across years. The results to date strongly suggest the potential for host-driven insect dynamics in this system.

If host plant resource allocation plays a significant role in herbivore dynamics, then identifying the factors that influence plant allocation becomes important for understanding the processes that ultimately drive variation in herbivore population densities. Relative allocation of meristems toward reproduction generally increases with cactus age, but it is unlikely that age can explain the variation in allocation strategies that we observed. The increases in RRE across years at the Sevilleta (fig. 2A) were greater than expected based on a 1-year increase in age (T. E. X. Miller, unpublished data). Across sites, mean plant age differed significantly between only two of our survey cactus populations (Aguirre Springs [8.9 years] vs. Sandia National Forest [11.0 years]; protected least significant difference test: \( P = .023 \)), and the younger population had greater RRE (fig. 2B). Winter precipitation could also be an important factor in tree cholla resource allocation because it is known to influence other components of cactus performance during the growing season (Bowers 1996; Pimienta-Barrios and del Castillo 2002). The observed increase in tree cholla RRE across years was associated with increases in winter precipitation in central New Mexico (New Mexico Climate Center, http://weather.nmsu.edu). However, Bowers (1996), studying another *Opuntia* species (*Opuntia engelmannii*) in the Sonoran Desert, found no relationship between winter precipitation and meristem allocation. Rather, she suggested that cacti maximize their fitness by alternating periods of high reproductive effort with periods of mostly vegetative growth, although no physiological mechanism for such switching has been identified (Bowers 1996). The interaction of genetic and en-

---

**Figure 1:** Projections of *Narnia* abundance (juvenile + adult) using the Akaike’s Information Criterion–selected model structure (table 1, model 13) and parameter estimates (table 2). Lines represent predicted abundances. Triangles represent data used to fit the model, and circles represent independent data. Each data set is split into low relative reproductive effort (RRE; dashed line, open symbols) and high RRE (solid line, filled symbols) according to the median RRE value of each data set (median values of model fitting and independent data sets were 0.36 and 0.37, respectively). Points represent averages of plants greater or less than the median. Predictions were based on average low RRE (0.17) and average high RRE (0.53).
The American Naturalist

Figure 2: A, Temporal variation in tree cholla relative reproductive effort (RRE) and cactus bug abundance at the Sevilleta. Tree cholla RRE increased significantly from 2003 to 2005 (variation along the X-axis: filled circles; variation along the Y-axis: open circles; $F = 5.39, df = 2,138, P < .0055$). This increase was associated with an increase in Narnia abundance that was greater for the late-season cohort (filled circles) than for the early-season cohort (open circles; $F = 10.67, df = 2,269, P < .0001$). Spatial variation in cactus RRE and Narnia abundance (correlation analysis using raw data points: $r = 0.42, P < .001, n = 124$). Numbers next to data points correspond to sites on map. Points in both figures are mean ± SE.

environmental factors in plant resource allocation continues to be an area of active research (Bazzaz and Grace 1997; Vuorisalo and Mutikainen 1999).

Finally, on the basis of the strength and consistency of the modeling and empirical results, we suggest that future work on the evolution and maintenance of plant life-history strategies should consider the ecological consequences of resource allocation for interactions with other trophic levels. Costs of reproduction are central to plant life-history theory (Harper 1977), and accurately quantifying these costs is important for understanding the selective advantages of alternative allocation strategies (Reekie 1999). It is well established that current reproduction in plants can have negative effects on future survival and fecundity, and theories of allocation trade-offs are based almost exclusively on these resource-based costs (Obeso 2002). However, a strictly resource-based perspective may be insufficient to predict and explain plant life-history strategies.

We found that patterns of resource allocation between life-history functions had consequences for the abundance and dynamics of plant antagonists. Increasing the relative investment of meristems in reproduction led to an increase in the frequency of interactions with specialist herbivores. Thus, in addition to resource-based costs, plant reproductive effort may also incur herbivore-mediated ecological costs. This component of reproductive costs is typically neglected (Klinkhamer et al. 1997; Holland et al. 2004), and the generality of a positive relationship between plant reproduction and susceptibility to herbivores remains unknown (Obeso 2002). Given that herbivores can reduce plant fitness (Crawley 1983, 1997; Doak 1992; Louda and Potvin 1995; Bigger and Marvier 1998; Maron 1998; Ehrlen 2003), explicit consideration of the ecological costs of reproduction is needed to refine our understanding of optimal life-history strategies in plants.

Acknowledgments

We gratefully acknowledge the faculty and staffs of the Sevilleta long-term ecological research (LTER) site, especially S. Collins and J. Johnson, and the Sevilleta National Wildlife Refuge, especially D. Pritchard and R. Robichaud, for providing logistical support and a stimulating research environment. B. Inouye provided helpful discussion on the theory and practice of the model fitting approach. The suggestions of J. Eckberg, E. Gering, K. Keeler, C. Mitra, J. Stiver, and two anonymous reviewers greatly improved this work and manuscript. Funding for this work was provided by grants to T.E.X.M. from the Sevilleta LTER (National Science Foundation grant DEB-0217774) and the University of Nebraska’s School of Biological Sciences and Initiative for Ecological and Evolutionary Analysis. T.E.X.M. was supported by a Graduate Assistance in Areas of National Need fellowship in environmental science from the U.S. Department of Education.

Literature Cited