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Stress, Coping, and Suicide Ideation in Chinese College Students

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Abstract

The study was to examine 1) whether stress and coping styles could significantly predict the probability of suicide ideation; 2) and whether coping styles were mediators or moderators on the association between life stress and suicide ideation. The survey was conducted in a sample of 671 Chinese college students. Approximately twenty percent students reported having suicide ideation. Life stress, active coping styles, and passive coping styles all had independent effect on the probability of suicide ideation. Passive coping styles, especially fantasizing, mediated the relation between life stress and suicide ideation. Moderation hypotheses were not supported. Implications of the findings and future directions were discussed.

Keywords: stress, coping, suicide ideation, mediator, moderator

Introduction

Suicide is the leading cause of death among 15–34 year olds in China, accounting for 19% of the deaths in this age group (Phillips, Li, & Zhang, 2002). Although the suicide rate among Chinese college students is not reported, a considerable number of suicide cases among this population have (Fan, 2006; Mooney, 2005). Suicide behavior has been conceptualized as a continuum from suicide ideation, to suicide attempt, to completed suicide (Lewinsohn et al., 1995). Suicide ideation is defined as thoughts of harming or killing oneself (Institute of Medicine, 2002). It is not only an important predictor of the suicide attempt and completed suicide but also a significant marker for other mental health problems among youth (see Thompson et al., in press). Thus, it is imperative to identify the characteristics of suicide ideation with the aim of developing effective prevention and intervention programs. The present study was aimed at investigating risk and protective factors of suicide ideation using a sample of Chinese college students.

The relationship between stress and suicide behaviors has been well-documented (e.g., Grover et al., 2009; Thompson et al., in press; Dixon et al., 1991). For example, Thompson et al. (in press) found that recent stressful life events, especially physical and psychological abuse, significantly predicted suicide ideation in youth, even after controlling for a host of other risk factors, including psychological distress. Compared to non-suicidal youth, suicidal psychiatric inpatients reported more stressful life events (Josepho and Plutchik, 1994; Wilson et al., 1995). The association between life stress and suicide risk has also been found in population-based college students (Dixon et al., 1991; Ügurlu and Ona, 2009).

Coping, as another important predictor, has also been extensively researched in understanding suicide behaviors (Wilson et al., 1995; see Speckens & Hawton, 2005, for a review). The strain theory of suicide argues that deficient coping, which is one of the most important sources of strain, may lead to suicidal behaviors (Zhang, Wieczorek,
Recruitment, procedure, and sample

This present study used the secondary survey data which had been collected by other Chinese colleagues from a comprehensive public university in South China in 2006. The survey was a part of the university effort to increase the understanding of the suicide stress, and coping, in order to enhance suicide prevention. The students were invited to voluntarily participate in an online survey through email. The survey was confidential and no identification information was linked to individual data. In the recruiting email, potential risks (e.g., temporary upset arising from questions about suicide behaviors), confidentiality, and procedures were explained prior to the participation. The survey link led participants to the secured survey website developed by the University Mental Health Center. By logging on the survey site and completing the survey, students gave their consent to participation. Students could freely withdraw at anytime of answering the online survey questions. This was a convenient sample. Total 790 invitation emails were sent out and the valid surveys were 689. The response rate was 87%. Eighteen cases aged older than 25 were excluded in the present analysis due to the age. Total 671 participants were included and male participants were 326 (48.6%). They were between the ages of 19 and
25 years (M = 21.8, SD = 1.65). There were freshman (N = 158, 23.5%), sophomore (N = 189, 28.2%), junior (N = 125, 18.6%), senior (N = 100, 14.9%), and first-year graduate students (N = 99, 14.8%) respectively. Twenty three percent of students were the only child in the family. Students left home at average 16.9 (SD = 5.2) years old. About four percent of them reported their parents (N = 29) were divorced. Around 21 percent students reported that their fathers’ educational levels were elementary school or middle school, 52 percent of them were high school or technique school, and around 36 percent of them were college, university or above the level. Around 41 percent students reported that their mothers’ educational levels were elementary school or middle school, around 32 percent of them were in high school or technique school, and 24 percent of them were college, university or above level. Before going to college, 43 percent (N = 288) participants lived in rural area and 48 percent (N = 322) in urban area respectively.

Measures

Coping was assessed by the Coping Style Questionnaire (CSQ; Xiao & Xu, 1996), which was developed by the Chinese scholars based on the coping inventories widely used in Western countries (e.g., Folkman and Lazarus, 1988; Stone and Neale, 1984). The items were created considering Chinese culture and language system. CSQ has accepted reliability and validity and has been widely used in coping research in Chinese college students (Cao et al., 2009; Xiao and Li, 1996; Xiao and Xu, 1996). The CSQ consisted of 62 items assessing 6 categories of coping efforts, including problem solving (e.g., “I try my best to change the situation and make it better.”), seeking help (e.g., “I will ask someone’s help to deal with the difficulties.”), rationalization (e.g., “when having conflict with someone, it is due to his or her bad personality.”), avoidance (e.g., “I do not think much about the situation which makes me upset”), fantasizing (e.g., “I often fantasize I am a superman and can deal with all the difficulties”), and self-blaming (e.g., “I often blame myself.”). The participants were asked whether or not they used the coping (No = 0), and If Yes, to rate on a 3-digit scale of the usefulness of this coping effort (not = 1, a little = 2, very = 3). Thus, each item was scored from 0 to 3. The total scores of each subscale were calculated. In addition, the total scores of problem solving and seeking help were summed to compose of the subscale of adaptive coping. The total scores of self-blaming, avoidance, and fantasizing were summed to compose of passive coping style. The Cronbach’s alphas for the six subscales were .73 (problem solving, 12 items), .63 (seeking help, 10 items), .66 (rationalizing, 11 items), .72 (avoidance, 10 items), .73 (fantasizing, 10 items), and .77 (self-blaming, 9 items). The Cronbach’s alphas of active coping style and passive coping style subscale were .79 and .87, respectively.

Life stress was measured using Adolescent Self-Rating Life Events Checklist (ASLEC, Liu, Liu, Yang, & Zhao, 1997). The ASLEC has accepted reliability and validity and has been successfully used in Chinese adolescents and young adults (Liu et al., 2000, 1997; Qiu and Yan, 2009). It consists of 6 subscales with total 26 items. The participants were asked whether or not the event happened in their life during the past 12 months (No = 0), and if Yes, to rate on a 5-digit scale of the stressfulness of this event (not stressful = 1 to extremely stressful = 5). Thus, each item was scored from 0 to 5. The total score of each subscale was calculated. The subscales were interpersonal relationship stress (6 items, e.g., “misunderstood or blamed wrongly by others”), school stress (5 items, e.g., “failing on exam or not getting the expected scores”), death and loss (3 items, e.g., “a relative or friend died”), physical health (3 items, e.g., “I got a severe illness”), setbacks (7 items, e.g., “I was fined.”) and others (2 items). The Cronbach’s alpha for these subscales are .79 (interpersonal relationship), .72 (school stress), .70 (death and loss), .86 (setbacks), .57 (physical health), and .52 (others), indicating an acceptable internal consistency of these measures. The overall stress was used in the analysis.

Suicide ideation was measured by one statement, “Have you thought about committing suicide?” There were five options, ranged from “Never = 1” to “Always = 5”.

Other controlled demographic variables included gender, age, if having siblings, and residence (rural or urban area). They were controlled in each model.

Data analysis

Logistic regression analysis

Suicide ideation was recoded as 0 (never = 1) and 1 (2 = once in a while to 5 = always). Life stress, active coping, and passive coping were all recoded into three levels (mean, one SD below, and one SD above), respectively. We tested if the overall stress, active coping and passive coping could significantly predict suicide ideation, after controlling for gender, age, if only child, and residence.

Moderation analyses

Active coping and passive coping were tested together as moderators of the relation between stress and suicide ideation using hierarchal linear regression analysis. In the first step, the predictor variables and control variables were entered into the model. In the second step, the interaction terms of life stress × active coping and life stress × passive coping were entered into the model. Mean-centered life stress and coping were used, in order to reduce collinearity between the interaction and main effect (Aiken & West, 1991). Missing data were dealt with listwise deletion.

Mediation analyses

We tested different coping styles as multiple mediators simultaneously in one single model. Preacher and Hayes (2008) listed four advantages of specifying and testing a single multiple mediation model, instead of separate simple
mediation models. First, it can determine whether an overall mediating effect of a set of variables exists, which can help researchers to conclude whether the set of variables mediates the effect of the independent variable on the dependent variable. Second, we can know the magnitude of one specific mediating effect, conditional on the presence of other mediators in the model. Third, the involvement of multiple mediators in one model can reduce the likelihood of parameter bias due to potential omitted variables. Fourth, it allows the researcher to contrast and determine the relative magnitude of each specific mediating effect.

Bootstrapping approach is suggested to assess whether the total and specific mediating effects were significant (Preacher & Hayes, 2008). The advantage of this approach is that multivariate normality of the sampling distributions of the total and specific mediating effects, which is rare in the real world, does not need to be assumed (Preacher & Hayes, 2008). To bootstrap the sampling distribution of the specific and total mediating effects, a sample size of \( n \) cases is replaced for the original sample. Using the new sample, the path coefficients will be re-estimated and the mediating effects can be calculated (Preacher & Hayes, 2008). In the present analysis, we repeated the process for 5000 times. Thus, 5000 estimates of the total and specific mediating effects were attained. The distributions of these 5000 estimates serve as empirical, nonparametric approximations of the sampling distributions of the total and specific mediating effects. The ninety-five percent bootstrap Confidence Intervals (CIs) of the total and mediating effects were reported. SPSS Macro Indirect software developed by Preacher and Hayes (2008) was used to conduct the analysis. It can be downloaded from http://www.fhayes.com/spss-sas-and-mplus-macros-and-code.html. Missing data were dealt with listwise deletion.

**Results**

**Means, standard deviation and correlation**

Approximately twenty percent students \((N = 137)\) reported having suicide ideation, of whom around four percent \((N = 5)\) students reported thinking of suicide often or always. Means, standard deviations and correlation coefficients among life stress, coping styles, and suicide ideation were seen in Table 1. As expected, suicide ideation was positively related to life stress and passive coping style and the three subscales of passive coping: self-blame, fantasizing, and avoidance. It was negatively associated with active coping style and the two subscales of active coping: problem solving and seeking help. Life stress was significantly positively related to passive coping and the three subscales of passive coping, while no significant correlations were found between life stress and active coping style and between life stress and the two subscales of active coping, respectively.

**Logistic regression analysis**

After controlling for gender, age, only child or not, and residence, life stress, active coping and passive coping significantly predict suicide ideation. The Odds Ratios (OR) were 2.02, .59, and 2.34, respectively. For example, when the life stress increased from mean to one standard deviation above the mean, the degree of suicide ideation would increase as twice as the degree of suicide ideation when life stress was at mean level. When the level of passive coping style increases from average to one standard deviation above, the degree of suicide ideation would also increase as about twice as the degree of suicide ideation when passive coping was at mean level. However, when active coping increased from mean to one standard deviation above, the degree of suicide ideation would decrease as close to half as the degree of suicide ideation when active coping was at mean level. The results were seen in Table 2.

**Mediation analyses**

First, active coping style and passive coping style were tested as multiple mediators for the relation between life stress and suicide ideation. The results showed that total mediating effect of active coping and passive coping was significant, and it was .0016, and 95% bootstrap CI [.0006, .0031]. The specific mediating effect of passive coping was significant. The effect was .0019, and 95% bootstrap CI [.0010 to .0035]. The specific mediating effect of active coping was at a significant trend. The effect was −.0003, and 95% bootstrap CI [−.0009, 0]. The mediating effect of passive

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<th>Table 1. Mean, standard deviation, and correlation matrix.</th>
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* \( p < .05 \); ** \( p < .01 \); *** \( p < .001 \)
coping was significantly greater than that of active coping. Active coping effect was −.0023 lower than passive coping, and 95% bootstrap CI [−.0040, −.0012]. The model with unstandardized path coefficients was seen in Figure 1.

Secondly, we tested each subscales of passive coping as multiple mediators of the relation between stress and suicide ideation. The results showed that only fantasizing was a significant mediator. The mediating effect of fantasizing was .0012, and 95% bootstrap CI [.0005, .0023]. The mediating effect of self-blaming was not significant, 95% bootstrap CI [−.0007, .0010]. The mediating effect of avoidance was not significant, either, 95% bootstrap CI [−.0003, .0012]. The model with unstandardized path coefficients was seen in Figure 2.

Thirdly, we tested the subscales of active coping as multiple mediators of the relation between life stress and suicide ideation. The specific mediating effect of seeking help and problem solving were both not significant, with 95% bootstrap CI [−.0006, .0001] and [−.0006, .0001], respectively.

**Moderation analyses**

The interaction term between life stress and active coping did not significantly predict suicide ideation, \( \beta = .03, t = .63, p > .05 \). Also, the interaction term between life stress and passive coping was not significant of predicting suicide ideation, either, \( \beta = .02, t = .37, p > .05 \).

**Discussion**

In our sample, we found that the prevalence rate of suicide ideation was close to 20 percent, which is close to the rates found in previous studies (Zhang and Jin, 1996; Zhang et al., 1999). It also fell in the range of 6%–39.2% reported in college students in other countries (Arria et al., 2009; Garlow et al., 2008; Mackenzie et al., 2011; Roh et al., 2007; Toprak et al., 2011). Because our sample came from only one university, generalizing the rate and comparing it with the findings of previous studies is difficult. A survey with a nationally representative sample of college students should be conducted.

One purpose of the present study was to examine whether life stress was significantly associated with suicide ideation. The finding that life stress significantly predicted the risk of suicide is consistent with previous findings in research targeted at Western (e.g., Thompson et al., in press; Uğurlu and Ona, 2009) and Chinese adolescents and young adults (e.g., Liu and Tein, 2005; Zhang et al., 2011). Life stress was an independent and unique predictor even after controlling for coping styles, which was also consistent with previous studies on college students (Dixon et al., 1991).
Some researchers have found common biological evidence between adverse life stress and suicidality (Sunnqvist, Westrin, & Träskman-Bendz, 2008). Thus, adverse life stress is a consistent and strong predictor in suicide behaviors. In China, college students are faced with a variety of sources of stress, including living far from home, a crowded living space, and strict regulations (Mooney, 2005), as well as parents' high academic expectations (Schneider & Lee, 1990). Furthermore, rapid social and economic changes also cause significant levels of stress for today's youth, for example, the fast pace of life, the unrealistically high price of housing, high unemployment, relatively low wages after graduation, and social unfairness. This means that a good education may not necessarily lead to a good living. Thus, all these stressors can be a cumulative risk for suicidal behaviors in college students. Researchers can investigate the effect of stress on suicide ideation from social ecological perspectives.

Consistent with previous studies on both Western and Chinese adolescents (e.g., Dixon et al., 1991; Lai Kwok and Shek, 2009), an active coping style was negatively associated with suicide ideation in the results of the current study. A lack of effective coping skills has been found in suicide victims (Zhang et al., 2010), suicidal inpatients (Wilson et al., 1995) and population-based adolescents and young adults with suicide ideation (Lai Kwok & Shek, 2009). A passive coping style was found to be positively associated with suicide ideation in the present study. It is consistent with previous studies that suicidal youths reported more maladaptive coping strategies than non-suicidal controls (Gould et al., 2004; Wilson et al., 1995). Many Chinese children today are the only child in the home and are treated as “little emperors.” They are overprotected by their parents. They are often ill-prepared to deal with stressful demands on university campuses when they go to college (Mooney, 2005). Thus, the suicide risk would be increased in these younger generations (see Mooney, 2005).

Another purpose of the present study was to examine the mediating and moderating role of coping style between life stress and suicide ideation. The results showed that the mediating role of active coping was significant at trend, and passive coping was a significant mediator between life stress and suicide ideation. Higher life stress increased use of passive coping, in turn, passive coping increased suicide ideation. Especially, fantasizing was a stronger mediator. It is understandable that students with fantasizing coping style may fantasize about the problems disappearing after committing suicide. The other mediators, avoidance and self-blame, were not significant. It might be due to the collinearity problem caused by high correlations among different passive coping styles (.59 -.68).

Unlike some previous studies (e.g., Clum and Febbraio, 1994; Grover et al., 2009; Yang and Clum, 1994), the hypotheses of the moderating roles of active coping and passive coping were not supported. According to the definitions of mediator and moderator, moderators are those preexisting and relatively stable characteristics of an individual that increase or decrease the effect of stress on psychological outcomes, while mediators represent the mechanisms that stress affects (Grant et al., 2003; Wadsworth et al., 2005). Researchers suggest that coping serves as a mediator in children and adolescents and as a moderator in adults, because coping is less stable in childhood and adolescence than in adulthood (Wadsworth et al., 2004, 2005). The results in the present study may indicate that college students’ coping is still developing and less stable, as children, early and middle adolescents do. As mentioned before, Chinese college students are usually overprotected by their parents, and it is always the first time for them to leave home when they go to college. They may begin to learn to cope in the real world during the college period. Thus, coping may serve as a mediator between life stress and suicide ideation.

Figure 2. The mediating model of self-blaming, fantasizing, and avoidance between life stress and suicide ideation.

* \( p < .05 \); ** \( p < .01 \); *** \( p < .001 \)
There are limitations in this study. First, it was a cross-sectional design, thus the directionality of the relationships among life stress, coping and suicide ideation was not determined. Second, stress, coping and suicide ideation were all measured by students’ self-reports, which may lead to shared method variance problems (Compas, Connor-Smith, Saltzman, Thomsen, & Wadsworth, 2001). The association between life stress, coping and suicide ideation could be spurious. In future studies, researchers can gather information from different sources, e.g., parents, siblings or peers. Third, suicide ideation was measured using one statement on a 5-point scale. It is better to use a suicide ideation scale with sound reliability and validity information (e.g., Watson, Goldney, Fisher, & Merritt, 2001). Finally, the sample was limited to one university. The interpretation of the findings has to be cautious for generalization.

Although the present study has these limitations, it has contributed to the current literature of the stress-coping framework in understanding suicidal behaviors using a Chinese sample. One of the contributions was produced by testing different coping styles as multiple mediators in one single model. To our knowledge, no researchers have tested coping styles using a multiple mediator model before. Furthermore, we also tested the moderating roles of active coping and passive coping styles in one single model, in order to compare with the mediation analysis. Finally, the mediation findings have implications for suicide prevention and intervention. Practitioners can develop program curricula aimed at reducing suicide ideation in college students by increasing the repertoires of adaptive coping strategies and decreasing their passive coping styles, especially the fantasizing coping style.

References


